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A new approach of ratio estimation in sample surveys

■ M. IQBAL JEELANI, S.E.H. RIZVI, MANISH KR. SHARMA, NAGEENA NAZIR AND FAHEEM JEELANI

See end of the paper for authors' affiliations

Correspondence to :

M IQBAL JEELANI Sher-e-Kashmir University of Agricultural Sciences and Technology of Jammu, JAMMU (J&K) INDIA Email : jeelani.miqbal@ gmail.com

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ABSTRACT : This article deals with the estimation of population mean under simple random sampling using a new form of ratio estimator. The expression for mean square error and bias has been obtained. An efficiency comparison is considered for proposed estimator with the classical ratio, product and exponential ratio estimator. Finally an empirical study is also carried out to judge the performance of proposed estimator.

KEY **WORDS** : Simple random sampling, Ratio estimator, Mean square error, Efficiency, AMS Classification: 62D05

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INTRODUCTION :

In sampling it is a well known fact that the proper use of auxiliary information supplied by auxiliary variable improves the efficiency of the estimators of population parameters of the study variable. Ratio estimation is one such example where this criterion is fulfilled. Ratio estimators are used for estimation of population mean when study variable and auxiliary variable are highly correlated with each other. Early historical developments of the ratio method of estimation are being well presented by, Sen (1993). Some of the modified ratio estimators given by Cochran (1977); Murthy (1967); Prasad (1989); Sen (1993); Singh and Tailor (2003 and 2005); Upadhyaya and Singh (1999); Yan and Tian (2010) and Jeelani and Maqbool (2013) are available in literature. of the study variable is defined as $\overline{y}_r = \overline{y}/_{\overline{x}} \overline{X} = \hat{R}\overline{X}$ assuming that population mean of the auxiliary variable is known \overline{y} and \overline{x} are sample means of study variable and auxiliary variable, and the mean square of the classical ratio estimator is $(MSE[\overline{y}_r) = \delta \overline{Y}^2[C_y^2 + C_x^2 - 2\rho_{yx}C_yC_x])$. Where $\delta = 1 - f'_n, \mathbf{f} = \frac{1}{n}, \hat{R}$ population ratio, ρ_{yx} is correlation coefficient between study variable and auxiliary variable, C_y and C_x are co-efficient of variations of study variable and auxiliary variable. Also $C_y = \frac{S_y}{\overline{Y}}, C_x = \frac{S_x}{X}, \quad yx = \frac{S_{yx}}{S_yS_x}$, respectively.

Suggested estimator:

In this section we suggest a new ratio estimator by

The classical ratio estimator for population mean \overline{Y}

adapting an exponential estimator given by, Bahl and Tuteja (1991). The suggested ratio estimator will take the form as given below;

$$\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{MIJ})} = \psi \mathbf{\bar{y}} + [\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{BT})} - \psi (\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{BT})})] = \psi \mathbf{\bar{y}} + (\mathbf{1} \cdot \psi) \mathbf{\bar{y}}_{\mathbf{E}(\mathbf{BT})} \quad (1)$$

where, $\mathbf{y}_{\mathbf{E}(\mathbf{BT})}$ is the exponential estimator proposed by, Bahl and Tuteja (1991) and is as follows;

$$\bar{\mathbf{y}}_{\mathbf{E}(\mathbf{BT})} = \bar{\mathbf{y}} \exp\left(\frac{\bar{\mathbf{X}} \cdot \bar{\mathbf{x}}}{\bar{\mathbf{X}} + \bar{\mathbf{x}}}\right)$$
(2)

Also in equ. (1) ψ is is a constant such that the mean square error of the proposed estimator $\bar{y}_{E(MLJ)}$ is minimized.

Estimation of bias and mean square error of the suggested estimator :

For calculation of bias and mean square error of the proposed estimator we have $\bar{\mathbf{y}} = \bar{\mathbf{Y}}(\mathbf{1}+\eta_0)$ and $\bar{\mathbf{x}} = \bar{\mathbf{X}}(\mathbf{1}+\eta_1)$, such that the expectation of η_0 and η_1 is equal to zero E $(\eta_0) = E(\eta_1) = 0$

Now we have

$$\mathbf{E}(_{0})^{2} = \left(\frac{\mathbf{S}_{y}}{\overline{\mathbf{Y}}}\right)^{2}$$
(3)

$$\mathbf{E}(_{1})^{2} = \left(\frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}}\right)^{2} \tag{4}$$

$$\mathbf{E}(\mathbf{0}, \mathbf{1}) = \frac{\mathbf{S}_{yx}}{\mathbf{S}_{y}\mathbf{S}_{x}} \left(\frac{\mathbf{S}_{y}}{\overline{\mathbf{Y}}} \mathbf{x} \frac{\mathbf{S}_{x}}{\overline{\mathbf{X}}}\right)$$
(5)

where,

$$\begin{split} \delta &= 1 - \frac{f_n}{n}, f = \frac{1}{n}, S_y = \sqrt{\frac{\sum_{i=1}^{N} (y_i - \overline{Y})^2}{N - 1}}, S_x = \sqrt{\frac{\sum_{i=1}^{N} (x_i - \overline{X})^2}{N - 1}}, \\ S_{yx} &= \sqrt{\frac{\sum (y_i - \overline{Y}) (x_i - \overline{X})}{N - 1}} \end{split}$$

Applying Taylor series expansion and expectations on equ. (1), we can have the Bias and MSE of the proposed estimator in the following steps;

$$y_{E(MIJ)} \cdot Y = Y\{ 0 - K - xK \}$$
(6)
where, $K = \frac{n_1}{2} - \frac{3}{8} \frac{1}{2} + \frac{0}{2} \frac{x}{2}$

Applying the value of K in equ. (6), the above equation changes to ;

$$\overline{\mathbf{y}}_{E(MIJ)} \cdot \overline{\mathbf{Y}} = \overline{\mathbf{Y}} \left\{ \begin{array}{ccc} \mathbf{n}_{1} & \mathbf{3} \\ \mathbf{0} \cdot \left(\frac{\mathbf{n}_{1}}{2} \cdot \frac{\mathbf{3}}{\mathbf{8}} & \mathbf{1}^{2} + \frac{\mathbf{0} \mathbf{x}}{2}\right) - \psi \left(\frac{\mathbf{n}_{1}}{2} \cdot \frac{\mathbf{3}}{\mathbf{8}} & \mathbf{1}^{2} + \frac{\mathbf{0} \mathbf{x}}{2}\right) \right\}$$
(7)

Utilizing the equ. (7), the Bias of proposed estimator will be;

$$Bais(y_{E(MLJ)}) = \delta \overline{Y} \left(\frac{S_x}{\overline{X}} \right)^2 [L - \psi L]$$
(8)

where,
$$L = \left(\frac{3}{8} \cdot \frac{Z_{yx}}{2}\right)$$

Taking the value of L in equ. (8), Bias of $y_{E(MIJ)}$ becomes;

Bais
$$(\mathbf{y}_{E(MIJ}) = \overline{\mathbf{Y}} \left(\frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}} \right)^2 \left[\left(\frac{3}{8} - \frac{\mathbf{Z}_{\mathbf{yx}}}{2} \right) - \left(\frac{3}{8} - \frac{\mathbf{Z}_{\mathbf{yx}}}{2} \right) \right]$$
(9)
where, $\mathbf{Z}_{\mathbf{yx}} = \frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}} \mathbf{S}_{\mathbf{x}}} / \left\{ \frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{Y}}} / \frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}} \right\}$

Now the MSE of $y_{E(MIJ)}$ will take the following form; $MSE(\bar{y}_{E(MIJ)}) - \bar{Y} = \bar{Y} \{_0 - \Gamma - x\Gamma \}$ (10) where

 $\Gamma = \frac{\eta_{l}}{2}$ Utilizing the value of Γ in equ. (10), we have;

$$\mathbf{MSE}(\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{MIJ})}) \cdot \mathbf{\bar{Y}} = \mathbf{\bar{Y}} \left\{ \begin{array}{cc} \mathbf{0} \cdot \frac{\eta_{1}}{2} & \frac{\eta_{1}}{2} \end{array} \right\}$$
(11)

Again by using Taylor series expansion and taking expectations and squaring on both sides in equ. (11) then the MSE of $\bar{y}_{E(ML)}$ will be

$$MSE(\overline{y}_{E(MIJ)}) = \overline{Y}\left\{ \left(\frac{S_x}{\overline{x}}\right)^2 + \frac{\{S_x/\overline{x}\}^2(1-)^2}{4} - \frac{S_{yx}}{S_y} \left(\frac{S_y}{\overline{y}}\right) \left(\frac{S_x}{\overline{x}}\right)(1-)\right\}$$
(12)

The optimum value of ψ to minimize the mean square error of equ. (1) can be easily found as;

$$\frac{\partial}{\partial} \mathbf{MSE}(\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{MIJ})}) = \mathbf{0}$$
(13)

Then

$$*=1-2\frac{\mathbf{S}_{\mathbf{y}\mathbf{x}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}}/\left\{\frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{y}}}/\frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{x}}}\right\}$$
(14)

where,

$$\mathbf{S}_{\mathbf{y}} = \sqrt{\frac{\sum_{i=1}^{N} (\mathbf{y}_{i} - \overline{\mathbf{Y}})^{2}}{N-1}}, \mathbf{S}_{\mathbf{x}} = \sqrt{\frac{\sum_{i=1}^{N} (\mathbf{x}_{i} - \overline{\mathbf{X}})^{2}}{N-1}},$$

$$\mathbf{S}_{\mathbf{yx}} = \sqrt{\frac{\Sigma(\mathbf{y}_{i} - \overline{\mathbf{Y}})(\mathbf{x}_{i} - \overline{\mathbf{X}})}{\mathbf{N} - \mathbf{1}}}$$

Internat. Res. J. Agric. Eco.& Stat., 8 (1) March, 2017 : 100-103 HIND AGRICULTURAL RESEARCH AND TRAINING INSTITUTE Now recalling the equ. (1) and applying ψ^* in it in place of ψ such that the MSE of $\overline{y}_{E(MIJ)}$ is minimum in place of.

Then the MSE ($y_{E(MIJ)}$) will be

$$\mathbf{MSE}(\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{MIJ})}) = \frac{\mathbf{1} \cdot \mathbf{f}}{\mathbf{n}} \mathbf{\overline{Y}}^{2} \left(\frac{\mathbf{S}_{\mathbf{y}}}{\mathbf{\overline{Y}}}\right)^{2} \left(\mathbf{1} - \left(\frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}}\right)^{2}\right)$$
(15)

Now;

$$\bar{\mathbf{y}}_{\mathbf{E}(\mathbf{MIJ})}^{*} = \left(1 - 2\frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}} / \left\{\frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{Y}}} / \frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}}\right\}\right) \bar{\mathbf{y}} + 2\frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}} / \left\{\frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{Y}}} / \frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}}\right\} \bar{\mathbf{y}}_{\mathbf{E}(\mathbf{BT})}$$
(16)

where, $\bar{\mathbf{y}}_{\mathbf{E}(\mathbf{BT})} = \bar{\mathbf{y}} \exp\left(\frac{\bar{\mathbf{X}} \cdot \bar{\mathbf{x}}}{\bar{\mathbf{X}} + \bar{\mathbf{x}}}\right)$

The mean square error of (15) will tend to mean square error of (16) by replacing ψ in (1) by ψ *, then the mean square error of (16) will be ;

$$\mathbf{MSE}(\mathbf{\bar{y}}_{\mathbf{E}(\mathbf{MIJ})}^{*}) = \frac{1-\mathbf{f}}{\mathbf{n}} \mathbf{\bar{Y}}^{2} \left(\frac{\mathbf{S}_{\mathbf{y}}}{\mathbf{\bar{Y}}}\right)^{2} \left(1 - \left(\frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}}\right)^{2}\right)$$
(17)

As we know the classical ratio, product and exponential estimator given by, Bahl and Tuteja (1991) for population mean \overline{Y} and there mean square errors by applying Taylor's first degree expansion are as follows;

$$(\bar{\mathbf{y}}_{\mathbf{r}}) = \frac{\bar{\mathbf{y}}}{\bar{\mathbf{x}}} \overline{\mathbf{X}} \text{ and } \mathbf{MSE}(\bar{\mathbf{y}}_{\mathbf{r}}) = \frac{1 \cdot \mathbf{f}}{n} \overline{\mathbf{Y}}^{2} \left\{ \left(\frac{\mathbf{S}_{\mathbf{y}}}{\bar{\mathbf{Y}}} \right)^{2} + \left(\frac{\mathbf{S}_{\mathbf{x}}}{\bar{\mathbf{X}}} \right)^{2} \left(1 \cdot 2 \frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}} \mathbf{S}_{\mathbf{x}}} / \left\{ \frac{\mathbf{S}_{\mathbf{y}}}{\bar{\mathbf{Y}}} / \frac{\mathbf{S}_{\mathbf{x}}}{\bar{\mathbf{X}}} \right\} \right\}$$
(18)
$$(\bar{\mathbf{y}}_{\mathbf{pr}}) = \bar{\mathbf{y}} \left(\frac{\bar{\mathbf{x}}}{\bar{\mathbf{X}}} \right) \text{ and } \mathbf{MSE}(\bar{\mathbf{y}}_{\mathbf{pr}}) = \frac{1 \cdot \mathbf{f}}{n} \overline{\mathbf{Y}}^{2} \left\{ \left(\frac{\mathbf{S}_{\mathbf{y}}}{\bar{\mathbf{Y}}} \right)^{2} + \left(\frac{\mathbf{S}_{\mathbf{x}}}{\bar{\mathbf{X}}} \right)^{2} \left(1 + 2 \frac{\mathbf{S}_{\mathbf{yx}}}{\mathbf{S}_{\mathbf{y}} \mathbf{S}_{\mathbf{x}}} / \left\{ \frac{\mathbf{S}_{\mathbf{y}}}{\bar{\mathbf{Y}}} / \frac{\mathbf{S}_{\mathbf{x}}}{\bar{\mathbf{X}}} \right\} \right\}$$
(19)

$$\overline{\mathbf{y}}_{\mathbf{E}(\mathbf{BT})} = \overline{\mathbf{y}} \exp\left(\frac{\overline{\mathbf{X}} \cdot \overline{\mathbf{x}}}{\overline{\mathbf{X}} + \overline{\mathbf{x}}}\right)$$
 and $\mathbf{MSE}(\overline{\mathbf{y}}_{\mathbf{E}(\mathbf{BT})}) = \frac{1 \cdot \mathbf{f}}{\mathbf{n}} \overline{\mathbf{Y}}^2$

$$\left\{ \left(\frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{Y}}} \right)^{2} + \left(\frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}} \right)^{2} / 4 - - \frac{\mathbf{S}_{\mathbf{y}\mathbf{x}}}{\mathbf{S}_{\mathbf{y}}\mathbf{S}_{\mathbf{x}}} \left\| \left(\frac{\mathbf{S}_{\mathbf{y}}}{\overline{\mathbf{Y}}} \right) \left(\frac{\mathbf{S}_{\mathbf{x}}}{\overline{\mathbf{X}}} \right) \right\| \right\}$$
(20)

For efficiency comparison we have;

The MSE of (17) will be less than MSE of (18), (19) and (20) if and only if

$$\psi < 3 - 4 \frac{S_{yx}}{S_y S_x} / \left\{ \frac{S_y}{\overline{Y}} / \frac{S_x}{\overline{X}} \right\}$$
 then $MSE(\overline{y}_{E(MIJ)} < MSE(\overline{y}_r)$
(21)

$$\psi < 3 + 4 \frac{S_{yx}}{S_y S_x} / \left\{ \frac{S_y}{Y} / \frac{S_x}{\overline{X}} \right\}$$
 then $MSE(\overline{y}_{E(MLI)} < MSE(\overline{y}_{pr})$
(22)

$$\psi < 2 - 4 \frac{S_{yx}}{S_y S_x} / \left\{ \frac{S_y}{\overline{Y}} / \frac{S_x}{\overline{X}} \right\} \text{ then } MSE(\overline{y}_{E(MIJ)}) < \overline{y}_{E(BT)}$$
(23)

Empirical study :

For empirical study two natural populations where considered from Singh and Chaudhary (1986) and Cochran (1997). The population parameters and the constants computed from the above populations are given Table 1. The efficiency comparison is given in Table 2.

Conclusion :

From theoretical discussion and numerical example, we infer that the proposed estimator is more efficient than the classical ratio, product estimator and the exponential ratio estimator given by, Bahl and Tuteja (1991), as the efficiency of the proposed estimator is much higher than the efficiencies of the classical ratio,

Table 1 : Parameters of the population											
Parameters Populations	N	n	\overline{Y}	x	ρ_{yx}	Sy	Cy	S _x	C _x		
Singh and Chaudhary (1986)	22	5	22.62	1467.54	0.92	33.04	1.46	2562.14	1.74		
Cochran (1997)	49	20	116.16	98.67	0.69	98.82	0.85	102.97	1.04		

Table 2 : Relative efficiency comparison

-	Estimators						
	(\overline{y}_r)	(\overline{y}_{pr})	$\overline{y}_{E(BT)}$	$\overline{y}_{E(MIJ)}$			
Singh and Chaudhary (1986)	180.63	167.13	210.13	286.41			
Cochran (1997)	128.27	110.56	166.32	212.19			

102 Internat. Res. J. Agric. Eco. & Stat., **8** (1) March, 2017 : 100-103 HIND AGRICULTURAL RESEARCH AND TRAINING INSTITUTE product estimator and the exponential ratio estimator given by, Bahl and Tuteja (1991). Hence, we strongly recommend that the proposed modified estimator may be preferred over the classical ratio, product estimator and the exponential ratio estimator given by, Bahl and Tuteja (1991) for the use of practical applications.

Authors' affiliations:

S.E.H. RIZVI, MANISH KR. SHARMA, NAGEENA NAZIR AND FAHEEM JEELANI, Sher-e-Kashmir University of Agricultural Sciences and Technology of Jammu, JAMMU (J&K) INDIA

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